

# The polytopes in a Poisson hyperplane tessellation

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## Abstract

For a stationary Poisson hyperplane tessellation  $X$  in  $\mathbb{R}^d$ , whose directional distribution satisfies some mild conditions (which hold in the isotropic case, for example), it was recently shown that with probability one every combinatorial type of a simple  $d$ -polytope is realized infinitely often by the polytopes of  $X$ . This result is strengthened here: with probability one, every such combinatorial type appears among the polytopes of  $X$  not only infinitely often, but with positive density.

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## 1 Introduction

Imagine a system  $\mathbf{H}$  of hyperplanes in Euclidean space  $\mathbb{R}^d$  ( $d \geq 2$ ) that induces a tessellation  $T_{\mathbf{H}}$  of  $\mathbb{R}^d$ . This means that any bounded subset of  $\mathbb{R}^d$  meets only finitely many hyperplanes of  $\mathbf{H}$  and that the components of  $\mathbb{R}^d \setminus \bigcup_{H \in \mathbf{H}} H$  are bounded. The closures of these components are then convex polytopes which cover  $\mathbb{R}^d$  and have pairwise no common interior points. The set of these polytopes is denoted by  $T_{\mathbf{H}}$ . We impose the additional assumption that the hyperplanes of  $\mathbf{H}$  are in general position; then each polytope of  $T_{\mathbf{H}}$  is simple, that is, each of its vertices is contained in precisely  $d$  facets. The polytopes appearing in  $T_{\mathbf{H}}$  may be rather boring; they could, for example, all be parallelepipeds. However, if the hyperplanes of  $\mathbf{H}$  have sufficiently many different directions, one can imagine that quite different shapes of polytopes appear in  $T_{\mathbf{H}}$ . Is it possible that every combinatorial type of a simple  $d$ -polytope is realized in  $T_{\mathbf{H}}$ ? This can be achieved in a much stronger sense.

In fact, suppose that  $\widehat{X}$  is a stationary and isotropic Poisson hyperplane process in  $\mathbb{R}^d$  (explanations are found in [7], for example). Its hyperplanes are almost surely in general position and induce a random tessellation of  $\mathbb{R}^d$ , denoted by  $X$ . The general character of the polytopes in  $X$  was recently investigated in [4]. For example, it was shown there that almost surely (a.s.) the translates of the polytopes in  $X$  are dense in the space of convex bodies in  $\mathbb{R}^d$  (with the Hausdorff metric). Another result was that a.s. the polytopes of  $X$  realize every combinatorial type of a simple  $d$ -polytope infinitely often. In the following, we improve the latter result considerably, replacing ‘infinitely often’ by ‘with positive density’. In the subsequent definition,  $B_n$  is the ball in  $\mathbb{R}^d$  with center at the origin and radius  $n \in \mathbb{N}$ , and  $\lambda_d$  denotes Lebesgue measure in  $\mathbb{R}^d$ . Further,  $\mathbb{1}_A$  is the indicator function of  $A$ .

**Definition 1.** *Let  $T$  be a tessellation of  $\mathbb{R}^d$ , and let  $A$  be a translation invariant set of polytopes in  $\mathbb{R}^d$ . We say that  $A$  **appears in  $T$  with density  $\delta$**  if*

$$\liminf_{n \rightarrow \infty} \frac{1}{\lambda_d(B_n)} \sum_{P \in T, P \subset B_n} \mathbb{1}_A(P) = \delta.$$

With this definition, we prove below that in a Poisson hyperplane tessellation in  $\mathbb{R}^d$  which is stationary and isotropic (that is, has a motion invariant distribution), almost surely every combinatorial type of a simple  $d$ -polytope appears with positive density. The actual result will, in fact, be more general: it is sufficient that the Poisson hyperplane tessellation is stationary and that its directional distribution, a measure on the unit sphere, is not zero on any nonempty open set and is zero on any great subsphere. The precise theorem is formulated in the next section.

## 2 Explanations

We work in the  $d$ -dimensional Euclidean space  $\mathbb{R}^d$  ( $d \geq 2$ ) with its usual scalar product  $\langle \cdot, \cdot \rangle$ . By  $\lambda_d$  we denote its Lebesgue measure, by  $o$  its origin, by  $B^d$  its unit ball (with  $nB^d =: B_n$ ), and by  $\mathbb{S}^{d-1}$  its unit sphere. The space of hyperplanes in  $\mathbb{R}^d$ , with its usual topology, is denoted by  $\mathcal{H}$ , and  $\mathcal{B}(\mathcal{H})$  is the  $\sigma$ -algebra of Borel sets in  $\mathcal{H}$ . Hyperplanes in  $\mathbb{R}^d$  are often written in the form

$$H(u, \tau) = \{x \in \mathbb{R}^d : \langle x, u \rangle = \tau\}$$

with  $u \in \mathbb{S}^{d-1}$  and  $\tau \in \mathbb{R}$ .

We assume that  $\widehat{X}$  is a stationary Poisson hyperplane process in  $\mathbb{R}^d$ , thus, a Poisson process in the space  $\mathcal{H}$  of hyperplanes, with the property that its distribution is invariant under translations (we refer, e.g., to [7] for more details). The *intensity measure*  $\widehat{\Theta}$  of  $\widehat{X}$  is defined by

$$\widehat{\Theta}(A) = \mathbb{E} \widehat{X}(A) \quad \text{for } A \in \mathcal{B}(\mathcal{H}).$$

Here  $\mathbb{E}$  denotes expectation, and we write  $(\Omega, \mathcal{A}, \mathbb{P})$  for the underlying probability space. It is assumed that  $\widehat{\Theta}$  is locally finite and not identically zero. That  $\widehat{X}$  is a Poisson process includes that

$$\mathbb{P}(\widehat{X}(A) = k) = e^{-\widehat{\Theta}(A)} \frac{\widehat{\Theta}(A)^k}{k!} \quad \text{for } k \in \mathbb{N}_0,$$

for any  $A \in \mathcal{B}(\mathcal{H})$  with  $\widehat{\Theta}(A) < \infty$ .

Since  $\widehat{X}$  is stationary, the measure  $\widehat{\Theta}$  has a decomposition

$$\widehat{\Theta}(A) = \widehat{\gamma} \int_{\mathbb{S}^{d-1}} \int_{-\infty}^{\infty} \mathbb{1}_A(H(u, \tau)) \, d\tau \varphi(du)$$

for  $A \in \mathcal{B}(\mathcal{H})$  (see [7], Theorem 4.4.2 and (4.33)). The number  $\widehat{\gamma} > 0$  is the *intensity* of  $\widehat{X}$ , and  $\varphi$  is a finite, even Borel measure on the unit sphere. It is called the *spherical directional distribution* of  $\widehat{X}$ . For any such measure  $\varphi$  and any number  $\widehat{\gamma} > 0$ , there exists a stationary Poisson hyperplane process in  $\mathbb{R}^d$  with these data, and it is unique up to stochastic equivalence.

The hyperplane process  $\widehat{X}$  induces a random tessellation of  $\mathbb{R}^d$ , which we denote by  $X$ . As usual, a random tessellation is formalized as a particle process; we refer again to [7].

Since we are considering only simple processes, it is convenient to identify such a process, which by definition is a counting measure, with its support, which is a locally finite set. In particular, a realization  $\widehat{X}(\omega)$  of  $\widehat{X}$  is also considered as a set of hyperplanes, and a realization of  $X$  is considered as a set of polytopes. The notations  $\widehat{X}(\omega)(\{H\}) = 1$  and  $H \in \widehat{X}(\omega)$  for a hyperplane  $H$ , for example, are therefore used synonymously.

The combinatorial type of a polytope  $P$  in  $\mathbb{R}^d$  is the set of all polytopes in  $\mathbb{R}^d$  that are combinatorially isomorphic to  $P$ . Now we can formulate our result.

**Theorem 1.** *Let  $X$  be a tessellation of  $\mathbb{R}^d$  that is induced by a stationary Poisson hyperplane process  $\tilde{X}$  with spherical directional distribution  $\varphi$ . Suppose that the support of  $\varphi$  is the whole unit sphere  $\mathbb{S}^{d-1}$  and that  $\varphi$  assigns measure zero to each great subsphere of  $\mathbb{S}^{d-1}$ . Then, with probability one, each combinatorial type of a simple  $d$ -polytope appears with positive density in  $X$ .*

Theorem 1 implies, trivially, that under its assumptions almost surely each combinatorial type of a simple  $d$ -polytope appears infinitely often in  $X$ . When the latter fact was proved, among other results, in [4], a tool was a strengthened version of the Borel–Cantelli lemma, due to Erdős and Rényi [3] (see also [5, p. 327]). When the note [4] was submitted, an anonymous referee wrote “that the use of ergodicity of the mosaic could lead to a possibly shorter alternative proof”, and he/she briefly indicated a possible approach. After thorough consideration, we preferred the more elementary Borel–Cantelli lemma. However, reconsideration revealed that ergodicity, applied in a different way, would lead to a stronger result, as far as the occurrence of combinatorial types is concerned. This is carried out in the following.

### 3 Proof

Let  $X$  satisfy the assumptions of Theorem 1. Under the only assumption that the spherical directional distribution of the stationary Poisson hyperplane tessellation  $X$  is zero on every great subsphere, it was shown in [7, Thm. 10.5.3] that  $X$  is mixing and hence ergodic. This requires a few explanations. To model  $X$  as a point process, we consider the space  $\mathcal{K}$  of convex bodies (nonempty, compact, convex subsets) in  $\mathbb{R}^d$  with the Hausdorff metric. By  $\mathcal{B}(\mathcal{K})$  we denote the  $\sigma$ -algebra of Borel sets in  $\mathcal{K}$ . Let  $\mathbf{N}_s(\mathcal{K})$  be the set of simple, locally finite counting measures on  $\mathcal{B}(\mathcal{K})$  and  $\mathcal{N}_s(\mathcal{K})$  its usual  $\sigma$ -algebra (for details see, e.g., [7, Sect. 3.1]). As underlying probability space  $(\Omega, \mathcal{A}, \mathbb{P})$ , on which  $X$  is defined, we can use  $(\mathbf{N}_s(\mathcal{K}), \mathcal{N}_s(\mathcal{K}), \mathbb{P}_X)$ , where  $\mathbb{P}_X$  is the distribution of  $X$ . For  $t \in \mathbb{R}^d$ , a bijective map  $\mathsf{T}_t : \eta \mapsto \mathsf{T}_t \eta$  of  $\mathbf{N}_s(\mathcal{K})$  onto itself is defined by

$$(\mathsf{T}_t \eta)(B) := \eta(B - t), \quad B \in \mathcal{B}(\mathcal{K}), \eta \in \mathbf{N}_s(\mathcal{K}).$$

Since  $X$  is stationary, we have

$$\mathbb{P}_X(\mathsf{T}_t A) = \mathbb{P}_X(A) \quad \text{for } A \in \mathcal{N}_s(\mathcal{K}),$$

thus  $\mathsf{T}_t$  induces a measure preserving map of  $\mathcal{N}_s(\mathcal{K})$  into itself. Let  $\mathcal{T} := \{\mathsf{T}_t : t \in \mathbb{R}^d\}$ . As shown in [7, Thm. 10.5.3], the dynamical system  $(\mathbf{N}_s(\mathcal{K}), \mathcal{N}_s(\mathcal{K}), \mathbb{P}_X, \mathcal{T})$  is mixing, that is,

$$\lim_{\|t\| \rightarrow \infty} \mathbb{P}_X(A \cap \mathsf{T}_t B) = \mathbb{P}_X(A) \mathbb{P}_X(B)$$

holds for all  $A, B \in \mathcal{N}_s(\mathcal{K})$ . It follows that the system is ergodic, which means that  $\mathbb{P}_X(A) \in \{0, 1\}$  for all  $A \in \mathbf{T} := \{A \in \mathcal{N}_s(\mathcal{K}) : \mathsf{T}_t A = A \text{ for all } t \in \mathbb{R}^d\}$ . Therefore, the ‘Individual Ergodic Theorem for  $d$ -dimensional Shifts’ yields the following.

**Proposition 1.** *Let  $f$  be an integrable random variable on  $(\mathbf{N}_s(\mathcal{K}), \mathcal{N}_s(\mathcal{K}), \mathbb{P}_X)$ . Then*

$$\lim_{n \rightarrow \infty} \frac{1}{\lambda_d(B_n)} \int_{B_n} f(\mathsf{T}_t \omega) \lambda(dt) = \mathbb{E} f$$

*holds for  $\mathbb{P}_X$ -almost all  $\omega \in \mathbf{N}_s(\mathcal{K})$ .*

We refer to Daley and Vere–Jones [2, Proposition 12.2.II] for a more general formulation (with hints to proofs of more general results in Tempel’man [6]). However, we have already incorporated into our Proposition 1 the information that in our case  $(\mathbf{N}_s(\mathcal{K}), \mathcal{N}_s(\mathcal{K}), \mathbb{P}_X, \mathcal{T})$  is ergodic, which yields that the limit is equal to the expectation of  $f$ .

We apply this Proposition in the following way. First we choose a center function  $c$  on  $\mathcal{K}$ ; for example, let  $c(K)$  denote the circumcenter of  $K \in \mathcal{K}$ , which is the center of the smallest ball containing  $K$ . Let  $A \in \mathcal{B}(\mathcal{K})$  be a translation invariant Borel set of convex bodies. Given any bounded Borel set  $B \in \mathcal{B}(\mathbb{R}^d)$ , we define

$$f(B, \omega) := \sum_{K \in X(\omega), c(K) \in B} \mathbb{1}_A(K)$$

for  $\omega \in \Omega$ , where we use  $(\Omega, \mathcal{A}, \mathbb{P}) = (\mathbf{N}_s(\mathcal{K}), \mathcal{N}_s(\mathcal{K}), \mathbb{P}_X)$  as the underlying probability space. Then  $f(B, \cdot)$  is measurable, and  $f(B + t, \omega) = f(B, \mathbb{T}_{-t}\omega)$  for  $t \in \mathbb{R}^d$ . The following generalizes an approach of Cowan [1] in the plane (“Tricks with small disks”). Assuming that  $n > 1$ , we have

$$\begin{aligned} & \int_{B_{n-1}} f(B_1 + t, \omega) \lambda_d(dt) \\ &= \sum_{K \in X(\omega)} \int_{\mathbb{R}^d} \mathbb{1}\{t \in B_{n-1}\} \mathbb{1}\{K \in A\} \mathbb{1}\{c(K) \in B_1 + t\} \lambda_d(dt). \end{aligned}$$

Since

$$\mathbb{1}\{t \in B_{n-1}\} \mathbb{1}\{c(K) \in B_1 + t\} \leq \mathbb{1}\{t \in -B_1 + c(K)\} \mathbb{1}\{c(K) \in B_n\},$$

we get

$$\begin{aligned} & \int_{B_{n-1}} f(B_1 + t, \omega) \lambda_d(dt) \\ & \leq \sum_{K \in X(\omega)} \int_{\mathbb{R}^d} \mathbb{1}\{t \in -B_1 + c(K)\} \mathbb{1}\{K \in A\} \mathbb{1}\{c(K) \in B_n\} \lambda_d(dt) \\ & = \lambda_d(B_1) f(B_n, \omega). \end{aligned}$$

Similarly,

$$\begin{aligned} & \int_{B_{n+1}} f(B_1 + t, \omega) \lambda_d(dt) \\ & \geq \sum_{K \in X(\omega)} \int_{\mathbb{R}^d} \mathbb{1}\{t \in -B_1 + c(K)\} \mathbb{1}\{K \in A\} \mathbb{1}\{c(K) \in B_n\} \lambda_d(dt) \\ & = \lambda_d(B_1) f(B_n, \omega). \end{aligned}$$

We conclude that

$$\begin{aligned} & \frac{\lambda_d(B_{n-1})}{\lambda_d(B_n)} \frac{1}{\lambda_d(B_{n-1})} \int_{B_{n-1}} f(B_1, \mathbb{T}_{-t}\omega) \lambda_d(dt) \\ & \leq \frac{\lambda_d(B_1)}{\lambda_d(B_n)} f(B_n, \omega) \\ & \leq \frac{\lambda_d(B_{n+1})}{\lambda_d(B_n)} \frac{1}{\lambda_d(B_{n+1})} \int_{B_{n+1}} f(B_1, \mathbb{T}_{-t}\omega) \lambda_d(dt). \end{aligned}$$

By Proposition 1, the lower and the upper bound converge, for  $n \rightarrow \infty$ , almost surely to  $\mathbb{E} f(B_1, \cdot)$ , hence a.s.

$$\lim_{n \rightarrow \infty} \frac{1}{\lambda_d(B_n)} f(B_n, \cdot) = \frac{\mathbb{E} f(B_1, \cdot)}{\lambda_d(B_1)}. \quad (1)$$

Now we assume in addition that there is a constant  $D > 0$  such that all convex bodies  $K \in A$  satisfy  $\text{diam } K \leq D$ , where  $\text{diam}$  denotes the diameter. The center function  $c$  satisfies  $c(K) \in K$ , hence if  $c(K) \in B_{n-D}$  (with  $n > D$ ) and  $\text{diam } K \leq D$ , then  $K \subset B_n$ . It follows that, for  $n > D$ ,

$$\begin{aligned} & \frac{\lambda_d(B_{n-D})}{\lambda_d(B_n)} \frac{1}{\lambda_d(B_{n-D})} \sum_{K \in X} \mathbb{1}_A(K) \mathbb{1}\{c(K) \in B_{n-D}\} \\ & \leq \frac{1}{\lambda_d(B_n)} \sum_{K \in X, K \subset B_n} \mathbb{1}_A(K) \\ & \leq \frac{1}{\lambda_d(B_n)} \sum_{K \in X} \mathbb{1}_A(K \in A) \mathbb{1}\{c(K) \in B_n\}. \end{aligned}$$

As  $n \rightarrow \infty$ , the lower and the upper bound converge a.s. to the right side of (1), hence a.s. we have

$$\delta(X, A) := \lim_{n \rightarrow \infty} \frac{1}{\lambda_d(B_n)} \sum_{K \in X, K \subset B_n} \mathbb{1}_A(K) = \frac{1}{\lambda_d(B^d)} \mathbb{E} \sum_{K \in X, c(K) \in B^d} \mathbb{1}_A(K). \quad (2)$$

Now we consider the special case where  $A_D$  is the set of polytopes that are combinatorially isomorphic to a given simple  $d$ -polytope  $P$  and have diameter at most  $D$ , for some fixed number  $D > 0$ . We remark that (2) shows that

$$\delta(X, A_D) = \frac{1}{\lambda_d(B^d)} \mathbb{E} \sum_{K \in X, c(K) \in B^d} \mathbb{1}\{K \in A_D\}, \quad (3)$$

It remains to show that

$$\mathbb{E} \sum_{K \in X, c(K) \in B^d} \mathbb{1}\{K \in A_D\} > 0. \quad (4)$$

For this, we use an argument from [4], which we recall for completeness.

Without loss of generality, we can assume that  $c(P) = o$ . Let  $F_1, \dots, F_m$  be the facets of  $P$ . We denote by  $B(x, \varepsilon)$  the ball with center  $x$  and radius  $\varepsilon > 0$ , set  $[B(x, \varepsilon)]_{\mathcal{H}} := \{H \in \mathcal{H} : H \cap B(x, \varepsilon) \neq \emptyset\}$ , and define

$$A_j(P, \varepsilon) := \bigcap_{v \in \text{vert } F_j} [B(v, \varepsilon)]_{\mathcal{H}}, \quad j = 1, \dots, m,$$

where  $\text{vert}$  denotes the set of vertices. Each hyperplane from  $A_j(P, \varepsilon)$  is said to be  $\varepsilon$ -close to  $F_j$ . A polytope  $Q$  is said to be  $\varepsilon$ -close to  $P$  if it has  $m$  facets  $G_1, \dots, G_m$  and, after suitable renumbering, the affine hull of  $G_j$  is  $\varepsilon$ -close to  $F_j$ , for  $j = 1, \dots, m$ . Since  $P$  is simple and  $c(P) = o$ , we can choose numbers  $D, \varepsilon_0 > 0$  such that for  $0 < \varepsilon \leq \varepsilon_0$ , the following is true:

- the sets  $A_1(P, \varepsilon), \dots, A_m(P, \varepsilon)$  are pairwise disjoint, and any hyperplanes  $H_j \in A_j(P, \varepsilon)$ ,  $j = 1, \dots, m$ , are the facet hyperplanes of a polytope  $Q$  that is  $\varepsilon$ -close to  $P$ .

- Any polytope  $Q$  that is  $\varepsilon$ -close to  $P$  satisfies the following:
- $Q$  is combinatorially isomorphic to  $P$ ,
- $Q \subset P + B^d$ ,
- $\text{diam } Q \leq D$ ,
- $c(Q) \in B^d$ .

That this can be achieved by suitable choices of  $D$  and  $\varepsilon_0$ , follows from easy continuity considerations and the fact that  $P$  is simple.

Now we define

$$C(P, \varepsilon) := \{H \in \mathcal{H} : H \cap (P + B^d) \neq \emptyset, H \notin A_j(P, \varepsilon) \text{ for } j = 1, \dots, m\}$$

and consider the event  $E(P, \varepsilon)$  defined by

$$\widehat{X}(A_j(P, \varepsilon)) = 1 \text{ for } j = 1, \dots, m \text{ and } \widehat{X}(C(P, \varepsilon)) = 0.$$

Let  $0 < \varepsilon \leq \varepsilon_0$ . The following was proved in [4]:

- If the event  $E(P, \varepsilon)$  occurs, then some polytope  $Q$  of the tessellation  $X$  is  $\varepsilon$ -close to  $P$  and hence satisfies  $Q \in A_D$  and  $c(Q) \in B^d$ ,
- The event  $\mathbb{P}(E(P, \varepsilon))$  has positive probability.

Now it follows that

$$\mathbb{E} \sum_{K \in X, c(K) \in B^d} \mathbb{1}\{K \in A_D\} \geq \mathbb{P}(E(P, \varepsilon)) > 0,$$

which proves (4).

The result is that  $\delta(X, A_D) > 0$  a.s. This implies, in particular, that with probability one the polytopes of the combinatorial type of  $P$  appear in  $X$  with positive density. Since there are only countably many combinatorial types, it also holds with probability one that each combinatorial type of a simple  $d$ -polytope appears in  $X$  with positive density.

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